

Lecture 11: Time in structured populations

MATH 242

11/6/2025

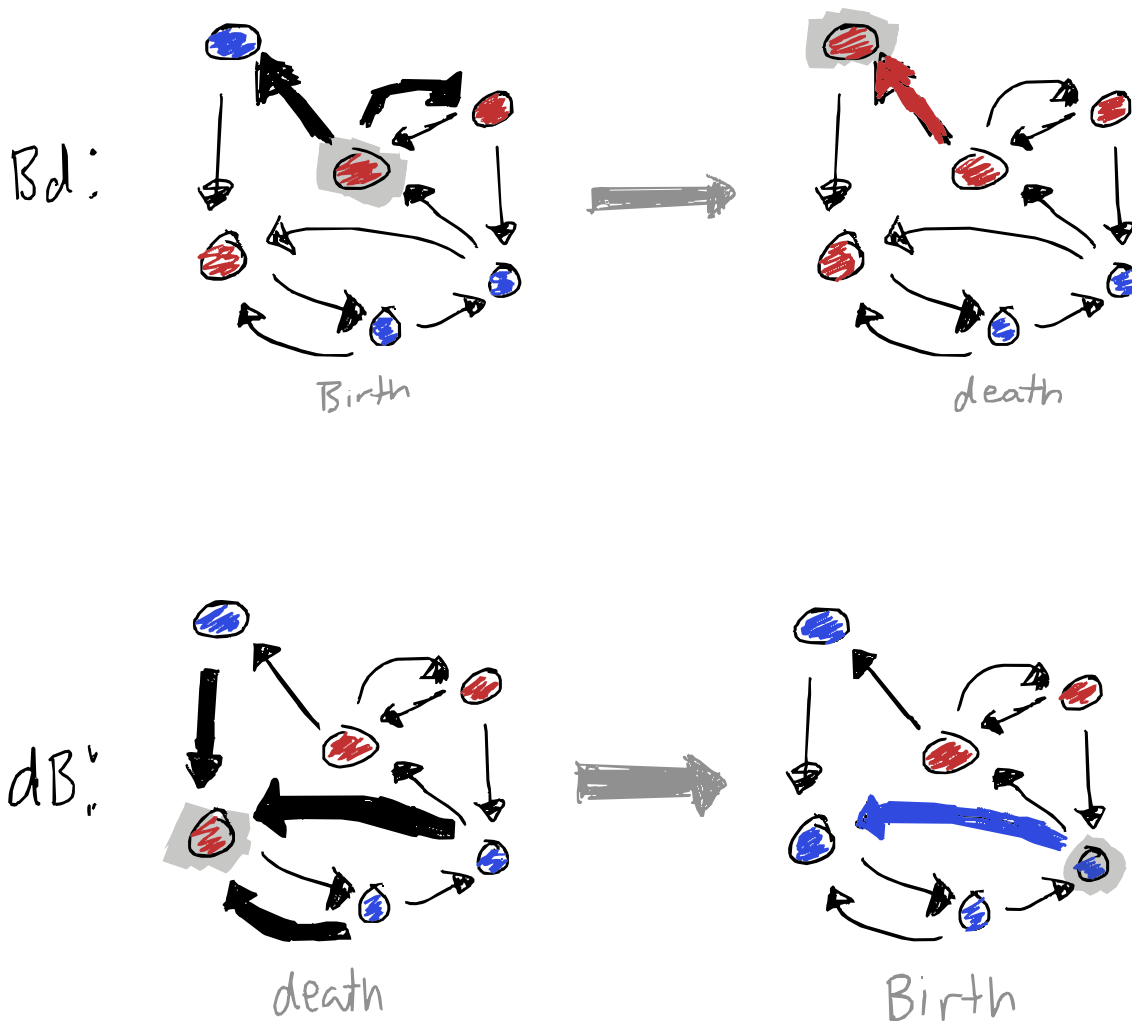
Last time, we saw in the Moran process:

(expected) absorption time is $\Theta(N \log N)$

N is population size

(expected) fixation time is $\Theta(N^2)$

Recall also Evolutionary Graph Theory as a way to model population structure



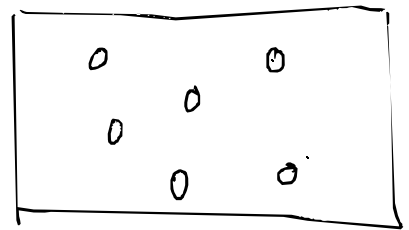
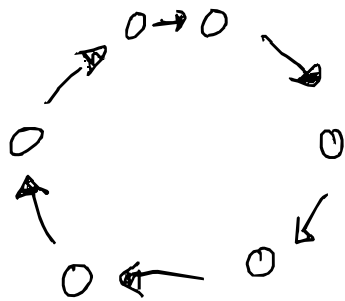
(note: B is capitalized since selection acts on Birth in these two dynamics)

Then the Moran process can be seen as a population with a complete unweighted graph w/ self loops as its structure (ie "well-mixed").

Isothermal theorem : Under Bd dynamics, if death

rate (i.e. "temperature") at each location is the same,
then fixation probability is the same as
Moran process = $\frac{1 - \frac{1}{r}}{1 - \frac{1}{r^N}}$ ^(or $\frac{1}{N}$ if $r=1$) for a single type **A**
individual w/ relative fitness r .

Ex: Directed cycle has same fixation probability
as Moran process in Bd dynamics



But what about fixation/absorption time?

For both directed cycles & well-mixed populations,
the # of type **A** in the population is sufficient
information to completely describe the state
(only true when directed cycle starts from a single type **A**)
(since births/deaths keep the types segregated)

Also, absorption/fixation times only depend on

$p_i^{\uparrow}$, $p_i^{\downarrow}$, and possibly φ_i^A ^{fixation prob}.

So even though $\chi_i = \frac{p_i^{\downarrow}}{p_i^{\uparrow}} = \frac{1}{r}$ for both directed cycle

and well-mixed, and fixation probabilities are the same by isothermal theorem, when $r=1$ we have (Bd)

well mixed $p_i^{\uparrow} = \frac{i}{N} \cdot \frac{N-i}{N} = \frac{i(N-i)}{N^2}$

$p_i^{\downarrow} = \frac{N-i}{N} \cdot \frac{i}{N} = \frac{i(N-i)}{N^2}$ $i=1, \dots, N-1$

$\Rightarrow p_i^{\circ} = 1 - (p_i^{\uparrow} + p_i^{\downarrow}) = 1 - \frac{2i(N-i)}{N^2}$

directed cycle $p_i^{\uparrow} = \frac{1}{N}, p_i^{\downarrow} = \frac{1}{N}$

$\Rightarrow p_i^{\circ} = 1 - \frac{2}{N}$

Since $p_i^{\circ, \text{directed}} > p_i^{\circ, \text{well-mixed}}$ for $i=1, \dots, N-1$,

The fixation/absorption times of directed cycles must be longer than that of well-mixed populations (in expectation).

How much longer??

From last lecture, $x_i \triangleq \mathbb{E}AT_i$

$$y_{i+1} \triangleq x_{i+1} - x_i = -\frac{1}{p_i^{\uparrow}} + \gamma_i y_i = -\left(\sum_{j=2}^i \frac{1}{p_j^{\uparrow}} \cdot \prod_{k=j+1}^i \gamma_k \right) + \prod_{k=1}^i \gamma_k$$

for $i=1, \dots, N-1$

$$x_1 = \frac{\sum_{i=0}^{N-1} r^{-i} \sum_{j=2}^i \frac{r^j}{p_j^{\uparrow}}}{\sum_{i=0}^{N-1} r^{-i}} \quad \text{for both directed cycles \& well-mixed}$$

For directed cycles, $p_j^{\uparrow} = \frac{1}{N}$ ($r=1$)

$$\text{So } (r=1) \\ X_1 = \frac{N \cdot \sum_{i=0}^{N-1} 1^{-i} \sum_{j=2}^i 1^j}{\sum_{i=0}^{N-1} 1^{-i}}$$

When $r=1$,

$$\mathbb{E} \bar{A}_2 = X_1 = \frac{N \cdot \sum_{i=0}^{N-1} \sum_{j=2}^i 1}{N} = \frac{1+2+3+\dots+(N-2)}{N} = \Theta(N^2).$$

For fixation times, recall

$$\begin{aligned} \mathbb{E} T_1 = X_1' &= \frac{1}{2p_1} \sum_{i=1}^{N-1} \prod_{j=2}^i q_j + \sum_{i=1}^{N-1} \sum_{j=3}^i \frac{1}{q_j} \prod_{k=j+1}^i q_k \\ &= \frac{N}{2} \sum_{i=1}^{N-1} \underbrace{\left(\prod_{j=2}^i \frac{1-\frac{1}{j}}{1+\frac{1}{j}} \right)}_{\leq 1} + N \cdot \sum_{i=1}^{N-1} \sum_{j=3}^i \underbrace{\left(\prod_{k=j+1}^i \frac{1-\frac{1}{k}}{1+\frac{1}{k}} \right)}_{\leq 1}. \end{aligned}$$

$$\begin{aligned} p_i &= p_i^* = \frac{1}{N} \\ q_i &= p_i \cdot (1 + \frac{1}{i}) \\ q_i &= \frac{1 - \frac{1}{i}}{1 + \frac{1}{i}} \end{aligned}$$

Thus

$$\begin{aligned} \mathbb{E} T_1 &\leq \frac{N^2}{2} + N \cdot \frac{N^2}{2} \\ &= O(N^3) \end{aligned}$$

and

$$\begin{aligned} \mathbb{E} T_1 &\geq N \cdot \sum_{i=1}^{N-1} \sum_{j=3}^i \frac{j(j+1)}{i(i+1)} \\ &\approx N \cdot \sum_{i=1}^{N-1} \frac{1}{\Theta(i^2)} = \Theta(i^3) \\ &\approx N \sum_{i=1}^{N-1} \Theta(i) \\ &= N \cdot \Theta(N^2) \\ &= \Omega(N^3). \end{aligned}$$

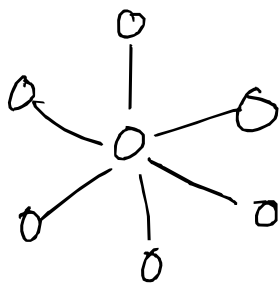
So, $EFT_1 = \Theta(N^3)$.

We have shown that there are simple 1D structures w/ same fixation prob. as Moran process but differing timescales.

Our analysis heavily uses the fact that the entire Markov chain state space is 1D.

But for a general population structure (eg graph), the dimensionality can become large.

Ex: Consider a star graph (undirected)



states can be reduced to

- (i) # leaves containing type **A**
- (ii) is center type **A**?

→ $\{0, 1, \dots, N-1\} \times \{\text{Yes, No}\}$.

This is a 2D representation of the state space and is sufficient information to compute fixation probs/times

However, the recurrence equations start to become more messy.

In general, we may need N dimensions to get exact expression for fixation probabilities/times.

When $r=1$, it is known that for undirected ^(connected) graphs $G=(V, E)$ under Bd dynamics

$$\begin{array}{l} \text{fixation prob of type } A \\ \text{when type } A \text{ occupies} \\ \text{locations } S \subseteq V \end{array} \stackrel{r=1}{=} \rho_{S,r=1}^A(G) = \frac{\sum_{u \in S} \frac{1}{\deg(u)}}{\sum_{u \in V} \frac{1}{\deg(u)}}$$

But for $r \neq 1$, there is no known exact, explicit, analytical, general formula known. :-

One approach to computationally approximate $\rho_{S,r}^A(G)$ for any r and S is the following:

(1) Run M independent Bd simulations starting from S

(2) Approximate $\rho_{S,r}^A(G)$ as $\hat{\rho}_{S,r}^A(G) = \frac{\# \text{ fixated simulations}}{M}$

How good is this approximation? (use a concentration bound)

$$\underbrace{\mathbb{P}\left\{ \left| \hat{\rho}_{S,r}^A(G) - \rho_{S,r}^A(G) \right| > \epsilon \right\}}_{\text{"bad approximation" event}} \stackrel{\substack{\text{Chernoff} \\ \text{bound}}}{\leq} \underbrace{2 \exp\left(-\frac{\epsilon^2 M}{4}\right)}_{\substack{\text{exponentially small} \\ \text{in } M}}$$

But in (1), we need the simulations to finish quickly or else our computational approach becomes intractable.

(arbitrary) (connected, unweighted)
For an undirected graph $G = (V, E)$,
it suffices to show that $\exists T: \mathbb{N} \rightarrow \mathbb{R}_{\geq 0}$ s.t.

$$\mathbb{E} AT_{s,r}(G) \leq T(N) \leq \text{poly}(N).$$

Then we can argue (by Markov inequality) that each one
of the M simulations will absorb within
time $100MT(N)$ w/ high prob;
(union bound)

$$\begin{aligned} \mathbb{P}\{\exists \text{ simulation longer than } 100MT(N)\} &\leq M \cdot \mathbb{P}\{\text{simulation takes longer than } 100MT(N)\} \\ &\leq M \frac{\mathbb{E} AT_{s,r}(G)}{100MT(N)} \\ &\leq \frac{1}{100}. \end{aligned}$$

So if we make M polynomial in N ,
our approximation has a very high chance
of being accurate and quick to compute. 😊

Q: How to show existence of $T(N)$?

Think of a game of chess (or any 2-player game)

At each point in the game, how do we evaluate
who is "winning" or "how quickly the game will end"?

Assumption: The more one player is winning,
the quicker the game will likely end.

More specifically, the rate at which a player

is taking the lead, the faster the game will likely end.

Q: How do we evaluate who is winning at a particular point in time?

A: We would like a function $\Phi: \Omega \rightarrow [0, 1]$ ^{state space}
s.t. $\exists$ disjoint subsets S_1, S_2 of Ω containing all absorbing states with:
(i) $\Phi(S_1) = 0 \quad \forall S_1 \subseteq S_1$ (ie. player 1 winning states)
(ii) $\Phi(S_2) = 1 \quad \forall S_2 \subseteq S_2$ (ie. player 2 winning states)
(iii) $\Phi(S') \geq \Phi(S) \Rightarrow$ it is more likely to absorb in S_2 than S_1 when starting at S' rather than S .

In chess, one can look k moves into the future and compute the number of player 1 wins vs player 2 wins.

But this is very complicated even if it is accurate.

A simpler method is to assign each piece a weight and calculate the weighted sum of pieces each player has.

This method is not as accurate but may suffice in many scenarios.

Lemma: If $X_0, X_1, X_2, \dots \in \Omega$ is a Markov chain, and $\tau \triangleq \min \{t : \Phi(X_t) = 1\}$, then

(i) $\mathbb{E}\tau < \infty$

$$(ii) \mathbb{E}[\Phi(X_{t+1}) - \Phi(X_t) | X_t] > \delta$$

$$\Rightarrow \mathbb{E} \tau < \frac{1-\mu}{\delta}$$

$$(iii) \Phi(X_0) \geq \mu$$

Now consider Bd process on connected, unweighted, undirected graph $G=(V, E)$, $r > 1$, except if type **A** goes extinct, restart the process.

Surely the absorption time of this augmented process is at least as long as the time for the typical, no-restart Bd process.

Weight each type **A** individual by the inverse of the degree of the vertex it occupies, so that

$$\Phi(S) = \frac{\sum_{u \in S} \frac{1}{\deg(u)}}{\sum_{u \in V} \frac{1}{\deg(u)}} = \rho_{S, r=1}^{\mathbf{A}}$$

$$\text{Then } \mathbb{E}[\Phi(X_{t+1}) - \Phi(X_t) | X_t = S] \quad \text{total fitness} = F(S) \triangleq r|S| + (N-|S|)$$

$$= \sum_{\substack{u \in S \\ v \notin S \\ (u,v) \in E}} \frac{r}{F(S)} \cdot \frac{1}{\deg(u)} \Phi(\{v\}) - \sum_{\substack{u \in S \\ v \notin S \\ (u,v) \in E}} \frac{1}{F(S)} \cdot \frac{1}{\deg(v)} \Phi(\{u\})$$

$$= \frac{r-1}{F(S) \cdot \sum_{u \in V} \frac{1}{\deg(u)}} \cdot \sum_{\substack{u \in S \\ v \notin S \\ (u,v) \in E}} \frac{1}{\deg(u) \deg(v)} \quad S \neq \emptyset, V$$

$$\geq \frac{r-1}{N(N-1)} \cdot \frac{1}{N(N-1)}$$

$$= \left(1 - \frac{1}{r}\right) N^{-4}.$$

Thus by the lemma,

$$\begin{aligned} \mathbb{E} A T_{x_0} &\leq \overbrace{(1-\mu)}^{\leq 1} \cdot \left(1 - \frac{1}{r}\right)^{-1} N^4 \\ &\leq \left(1 - \frac{1}{r}\right)^{-1} N^4 =: T(N). \end{aligned}$$

$\Rightarrow$ fast algorithm for approximating $\varphi_{s,r \geq 1}^A$ for any undirected, unweighted graph.

What about directed graphs?

Unfortunately, it is known to be likely intractable in the worst case (#P-hard). 😞

But for certain graphs (eg. regular, "balanced", etc) the simulation approach still works.

Some directed graphs take $\sim \exp(N)$ time to absorb.



```
In [2]: import numpy as np
import matplotlib.pyplot as plt
import seaborn as sns
import pandas as pd
import random
import enum
from dataclasses import dataclass
```

```
In [4]: class Type(enum.Enum):
    A = 0
    B = 1

@dataclass(frozen=True)
class SimulationResult:
    N: int
    population_structure: str
    fixation: bool
    time: int
```

```
In [5]: def simulate_well_mixed(N: int) -> SimulationResult:
    i = 1 # number of type A individuals
    t = 0 # time
    while 0 < i < N:
        birther_type, = random.choices([Type.A, Type.B], weights=(i, N-i))
        dier_type, = random.choices([Type.A, Type.B], weights=(i, N-i))
        if birther_type == Type.A and dier_type == Type.B:
            i += 1
        if birther_type == Type.B and dier_type == Type.A:
            i -= 1
        t += 1
    return SimulationResult(
        N=N,
        population_structure='well-mixed',
        fixation=i==N,
        time=t,
    )

def simulate_directed_cycle(N: int) -> SimulationResult:
    i = 1 # number of type A individuals
    t = 0 # time
    while 0 < i < N:
        birther_type, = random.choices([Type.A, Type.B], weights=(i, N-i))
        dier_type, = random.choices(
            [Type.A, Type.B],
            weights=(i-1, 1) if birther_type == Type.A else (1, N-i-1))
        if birther_type == Type.A and dier_type == Type.B:
            i += 1
        if birther_type == Type.B and dier_type == Type.A:
            i -= 1
        t += 1
    return SimulationResult(
        N=N,
        population_structure='directed-cycle',
```

```

        fixation=i==N,
        time=t,
    )

def simulate(N: int, population_structure: str) -> SimulationResult:
    if population_structure == 'well-mixed':
        return simulate_well_mixed(N)
    elif population_structure == 'directed-cycle':
        return simulate_directed_cycle(N)
    else:
        raise ValueError(f'Unknown population structure: {population_structure}')

```

```

In [6]: N = 100
        TRIALS = 1000
        data = [
            simulate(n, population_structure)
            for population_structure in ('well-mixed', 'directed-cycle')
            for n in range(1, N+1)
            for _ in range(TRIALS)
        ]
        df = pd.DataFrame(data)

```

```

In [7]: df.sample(n=10)

```

```

Out[7]:

```

	N	population_structure	fixation	time
97178	98	well-mixed	False	67
182155	83	directed-cycle	True	49925
117962	18	directed-cycle	True	955
37218	38	well-mixed	False	2548
5722	6	well-mixed	False	5
69284	70	well-mixed	False	236
185556	86	directed-cycle	False	92
164610	65	directed-cycle	False	34
97390	98	well-mixed	False	165
86300	87	well-mixed	False	44

```

In [11]: with sns.plotting_context(context='talk'):
        g = sns.lineplot(
            data=df,
            x='N',
            y='fixation',
            hue='population_structure',
        )

        Ns = np.arange(1, N)

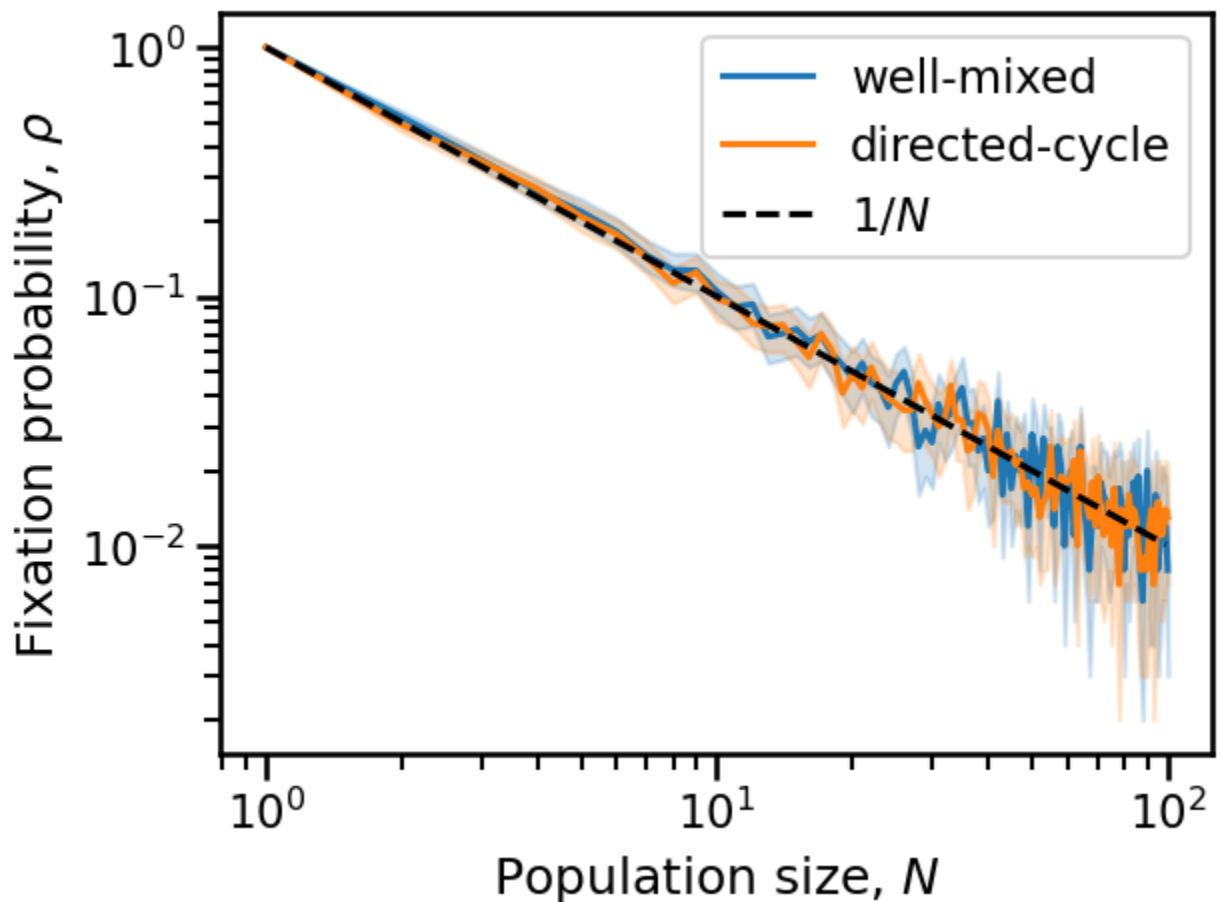
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```

sns.lineplot(
    x=Ns,
    y=1/Ns,
    label='$1/N$',
    ax=g,
    color='black',
    linestyle='--',
)

_ = g.set(
    xscale='log',
    yscale='log',
    xlabel='Population size, $N$',
    ylabel='Fixation probability, $\rho$',
)

```



```

In [13]: with sns.plotting_context(context='talk'):
    g = sns.lineplot(
        data=df,
        x='N',
        y='time',
        hue='population_structure',
    )

    Ns = np.arange(1, N)

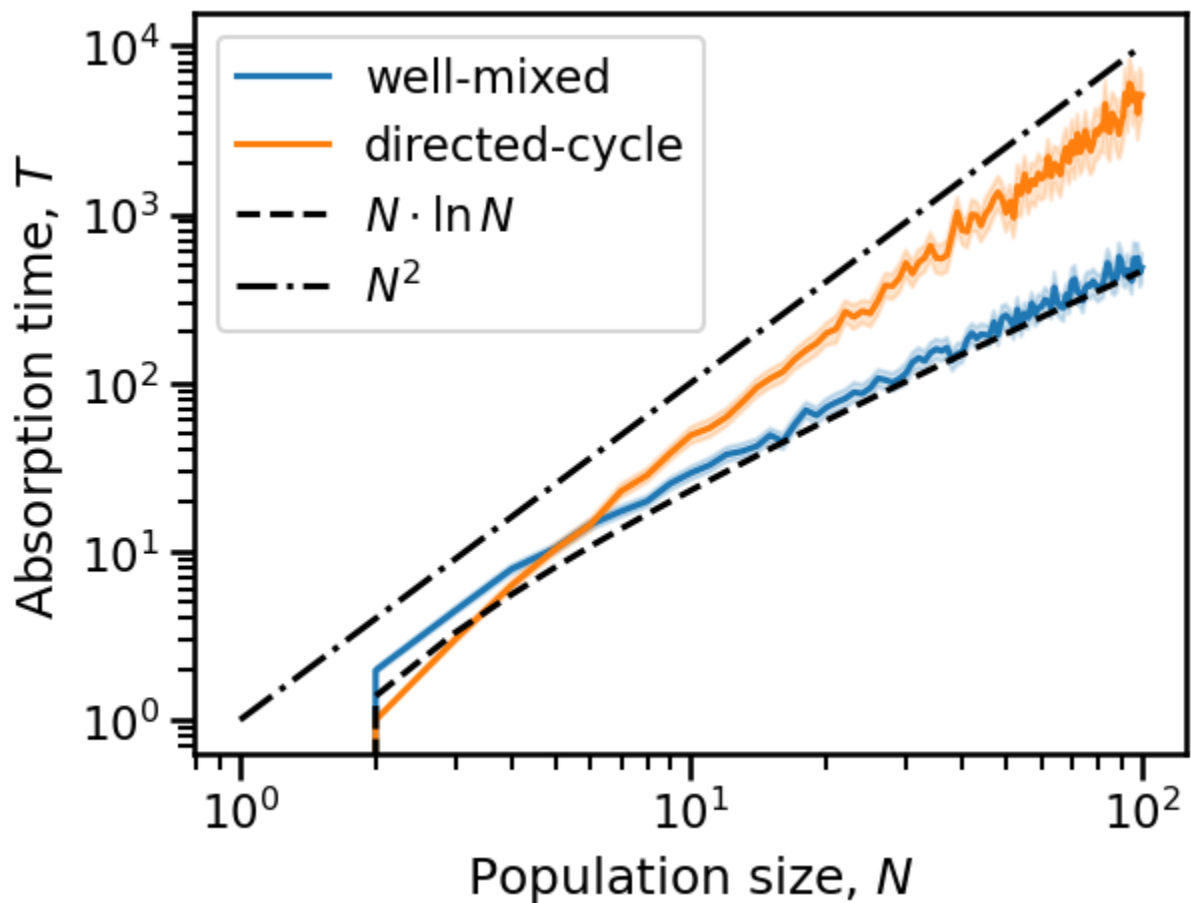
```

```

sns.lineplot(
    x=Ns,
    y=Ns*np.log(Ns),
    label='$N \cdot \ln N$',
    ax=g,
    color='black',
    linestyle='--',
)
sns.lineplot(
    x=Ns,
    y=Ns**2,
    label='$N^2$',
    ax=g,
    color='black',
    linestyle='-.',
)

- = g.set(
    xscale='log',
    yscale='log',
    xlabel='Population size, $N$',
    ylabel='Absorption time, $T$',
)

```



```
In [14]: with sns.plotting_context(context='talk'):
```

```

g = sns.lineplot(
    data=df[df['fixation']==True],
    x='N',
    y='time',
    hue='population_structure',
)

Ns = np.arange(1, N)
sns.lineplot(
    x=Ns,
    y=Ns**2,
    label='$N^2$',
    ax=g,
    color='black',
    linestyle='--',
)
sns.lineplot(
    x=Ns,
    y=Ns**3,
    label='$N^3$',
    ax=g,
    color='black',
    linestyle='-.',
)

_ = g.set(
    xscale='log',
    yscale='log',
    xlabel='Population size, $N$',
    ylabel='Fixation time, $T$',
)

```

